

Introduction to ctsmr

(Based on slides created by Rune Juhl)

Peder Bacher

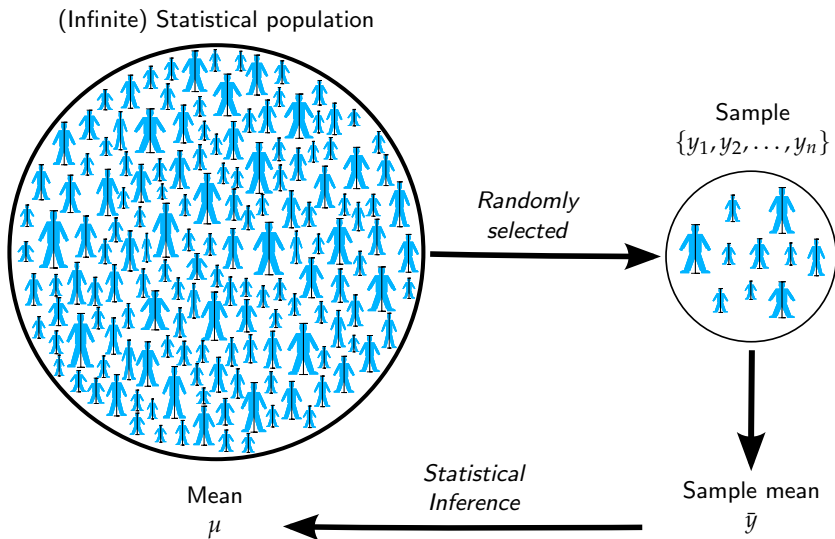
DTU Compute, Dynamical Systems
Building 303B, Room 010
DTU - Technical University of Denmark
2800 Lyngby – Denmark
e-mail: pbac@dtu.dk

Summer school 2023:

Time Series Analysis - with a focus on Modelling and Forecasting in Energy Systems

Overview

Population and sample



Parameter estimation with example

Simplest example: a constant model for the mean

- Model

$$Y_i = \mu + \varepsilon_i \quad , \text{ where } \varepsilon_i \sim N(0, \sigma^2) \text{ and i.i.d.}$$

- i.i.d.: identically and independent distributed
- The parameters are: the mean μ and the standard deviation σ
- We take a sample of $n = 100$ observations

$$(y_1, y_2, \dots, y_{100})$$

Likelihood

The likelihood is defined by the joint probability function of the data

$$L(\mu, \sigma) \equiv p(y_1, y_2, \dots, y_{100} | \mu, \sigma)$$

Hence, it's a function of the two parameters (the sample is observed, so it is not varying).

Due to independence

$$= \prod_{i=1}^{100} p(y_i | \mu, \sigma)$$

In our model the error $\varepsilon_i = Y_i - \mu$ is normal distributed (Gaussian), so

$$p(y_i | \mu, \sigma) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(y_i - \mu)^2}{2\sigma^2}\right) \quad (1)$$

Maximum likelihood estimation

Parameter estimation

$$\hat{\theta} = \arg \min_{\theta \in \Theta} \left(-\ln(L(\theta)) \right)$$

where $\theta = (\mu, \sigma)$

Maximum likelihood estimation

Parameter estimation

$$\hat{\theta} = \arg \min_{\theta \in \Theta} \left(-\ln(L(\theta)) \right)$$

where $\theta = (\mu, \sigma)$

Run the example in R

Likelihood for time correlated data

Given a sequence of measurements \mathcal{Y}_N

$$\begin{aligned} L(\theta) &= p(\mathcal{Y}_N|\theta) = p(y_N, y_{N-1}, \dots, y_0|\theta) \\ &= \left(\prod_{k=1}^N p(y_k|\mathcal{Y}_{k-1}, \theta) \right) p(y_0|\theta) \end{aligned}$$

Parameter estimation

$$\hat{\theta} = \arg \min_{\theta \in \Theta} (-\ln(L(\theta)))$$

Likelihood for time correlated data

If Gaussian

$$\hat{y}_{k|k-1} = E[y_k | \mathcal{Y}_{k-1}, \theta]$$

$$R_{k|k-1} = V[y_k | \mathcal{Y}_{k-1}, \theta]$$

$$\varepsilon_k = y_k - \hat{y}_{k|k-1}$$

then the likelihood is

$$L(\theta) = \left(\prod_{k=1}^N \frac{\exp(-\frac{1}{2} \varepsilon_k^T R_{k|k-1}^{-1} \varepsilon_k)}{\sqrt{|R_{k|k-1}|} \sqrt{2\pi}^l} \right)$$

Maximised using quasi Newton (in practise always minimize the negative log-likelihood)

Kalman filter

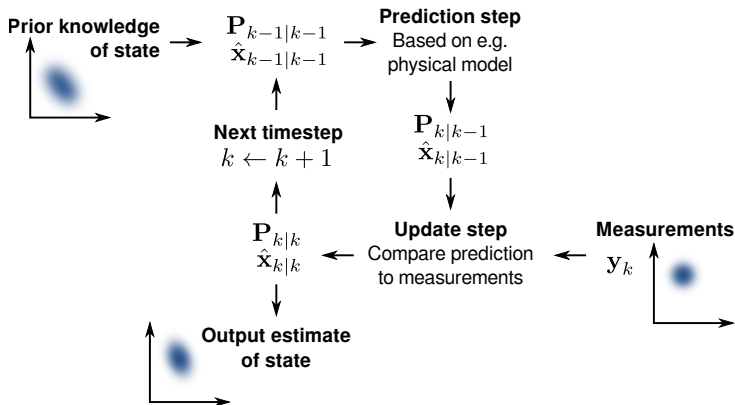


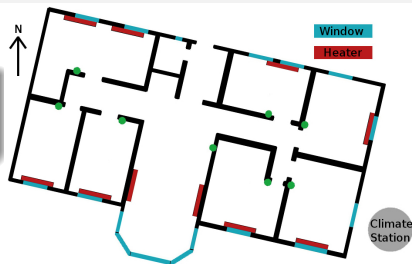
Figure: "Basic concept of Kalman filtering" by Petteri Aimonen. Wikipedia

Example: Selecting **a suitable grey-box model** for the heat dynamics of a building

Test case: One floored 120 m² building

Objective

Find the best model describing the heat dynamics of this building



Data

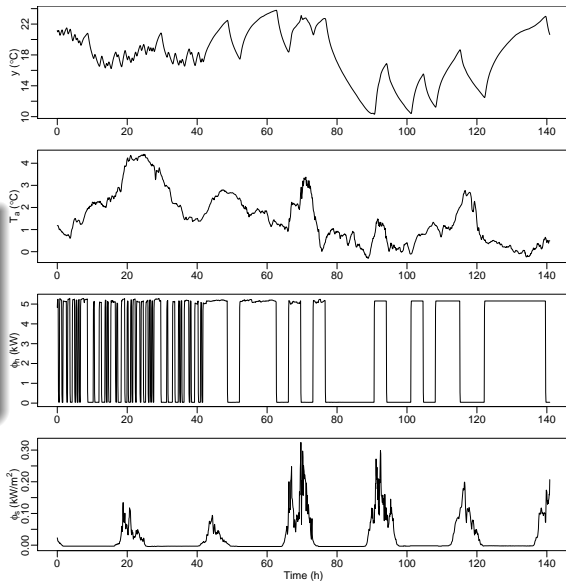
Measurements of:

y_t Indoor air temperature

T_a Ambient temperature

Φ_h Heat input

Φ_s Global irradiance



Two big challenges when modelling with data

- **Model selection:** How to decide which model is most appropriate to use?
We are looking for a model which gives us un-biased estimates of physical parameters of the system. This requires that the applied model is neither too simple nor too complex
- **Model validation:** How to validate the performance of a dynamical model?
We need to assess if the applied model fulfill assumptions of white-noise errors, i.e. that the errors show no lag-dependence

Model selection

Likelihood ratio test: Test for model expansion

Say we have a model and like to find out if an expanded version will give a significantly better description of data i.e. give an answer to: Should we use the expanded model instead of the one we have?

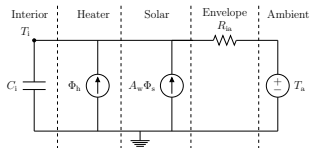
The likelihood ratio test

$$\lambda(\mathbf{y}) = -2 \ln \left[\frac{L_{\text{sub}}(\hat{\boldsymbol{\theta}}_{\text{mle,sub}})}{L(\hat{\boldsymbol{\theta}}_{\text{mle}})} \right]$$

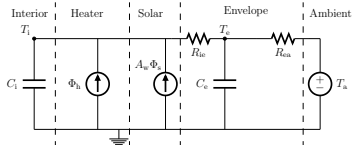
can be applied to test for significant improvement of the expanded model (with maximum likelihood $L(\hat{\boldsymbol{\theta}}_{\text{mle}})$) over the sub-model (with maximum likelihood $L_{\text{sub}}(\hat{\boldsymbol{\theta}}_{\text{mle,sub}})$)

Test for expansion

Simplest model

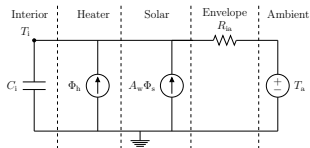


First extension: building envelope part ($T_i T_e$)

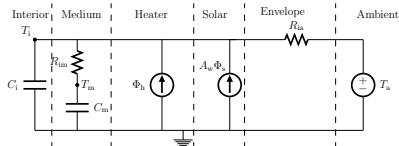


Test for expansion

Simplest model

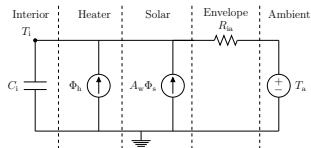


First extension: indoor medium part ($T_i T_m$)

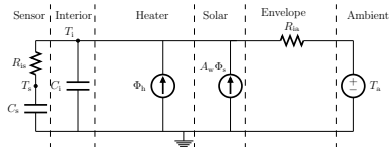


Test for expansion

Simplest model

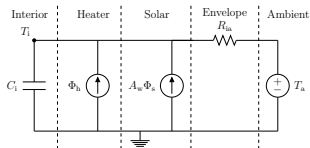


First extension: sensor part ($T_i T_s$)

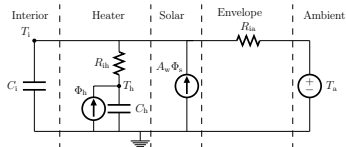


Test for expansion

Simplest model

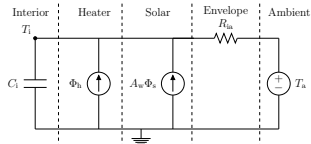


First extension: heater part (T_iTh)



Test for expansion

Simplest model



First extension: Which one??

$T_i T_e$, $T_i T_m$, $T_i T_s$, or $T_i T_h$?

Log-likelihoods

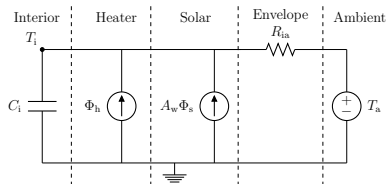
Simplest	<i>Ti</i>			
$l(\theta; \mathcal{Y}_N)$	2482.6			
<i>m</i>	6			
<hr/>				
Expanded	<i>TiTe</i>	<i>TiTm</i>	<i>TiT_s</i>	<i>TiTh</i>
$l(\theta; \mathcal{Y}_N)$	3628.0	3639.4	3884.4	3911.1
<i>m</i>	10	10	10	10

Likelihood-ratio test

Sub-model	Model	$m - r$	p-value
<i>Ti</i>	<i>TiTh</i>	4	$< 10^{-16}$

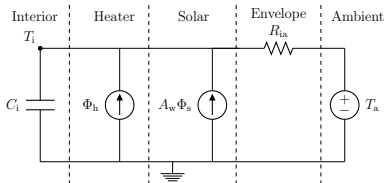
Identify the best physical model for the data

Simplest model

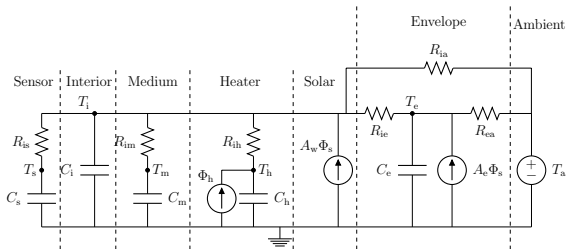


Identify the best physical model for the data

Simplest model

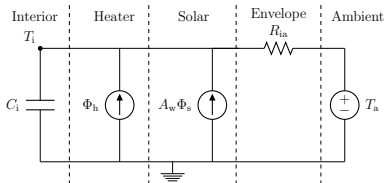


Most complex model applied



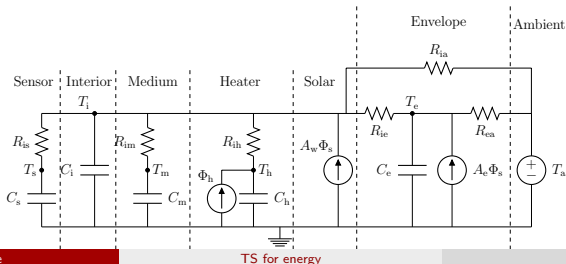
Identify the best physical model for the data

Simplest model



The best model for the given data is probably in between

Most complex model applied



Iteration	Models			
Start $l(\theta; \mathcal{Y}_N)$ m	Ti 2482.6 6			
1	$TiTe$ 3628.0 10	$TiTm$ 3639.4 10	$TiTs$ 3884.4 10	$TiTh$ 3911.1 10
2	$TiThTs$ 4017.0 14	$TiTmTh$ 5513.1 14	$TiTeTh$ 5517.1 14	
3	$TiTeThRia$ 5517.3 15	$TiTeThAe$ 5520.5 15	$TiTmTeTh$ 5534.5 18	$TiTeThTs$ 5612.4 18
4	$TiTeThTsRia$ 5612.5 19	$TiTmTeThTs$ 5612.9 22	$TiTeThTsAe$ 5614.6 19	
5	$TiTmTeThTsAe$ 5614.6 23	$TiTeThTsAeRia$ 5614.7 20		

Iteration	Sub-model	Model	$m - r$	$-2\log(\lambda(y))$	p-value
1	<i>Ti</i>	<i>TiTh</i>	4	4121	$< 10^{-16}$
2	<i>TiTh</i>	<i>TiTeTh</i>	4	4634	$< 10^{-16}$
3	<i>TiTeTh</i>	<i>TiTeThTs</i>	4	274	$< 10^{-16}$
4	<i>TiTeThTs</i>	<i>TiTeThTsAe</i>	1	6.4	0.011
5	<i>TiTeThTsAe</i>	<i>TiTeThTsAeRia</i>	1	0.17	0.68

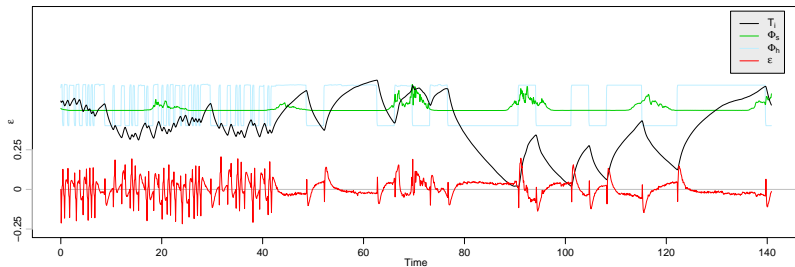
Model validation

How can the performance of a dynamical model be evaluated?

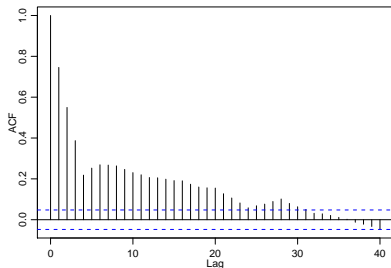
- We assume that the residuals are i.i.d and normal
- Auto-Correlation Function (ACF) and Cumulated Periodogram (CP) of the errors are the basic tools
- Time series plots of the inputs, outputs, and the errors are valuable for pointing out model deficiencies

Evaluate the simplest model

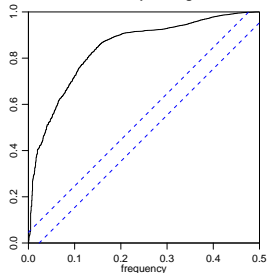
Inputs and residuals



ACF of residuals

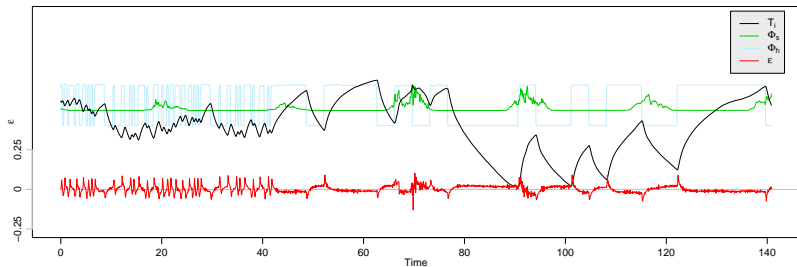


Cumulated periodogram

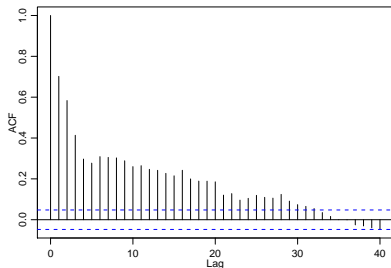


Evaluate the model selected in step one

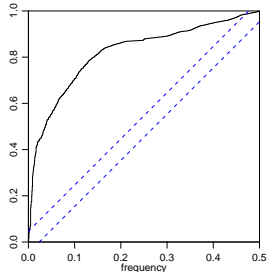
Inputs and residuals



ACF of residuals

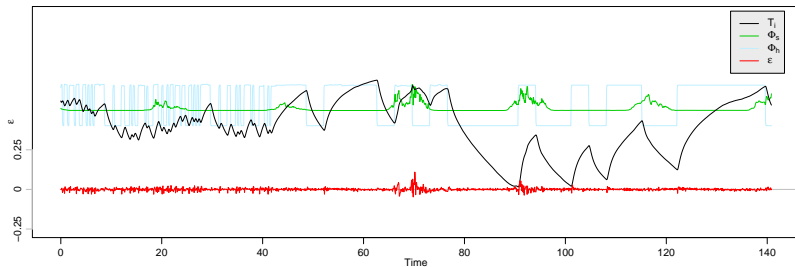


Cumulated periodogram

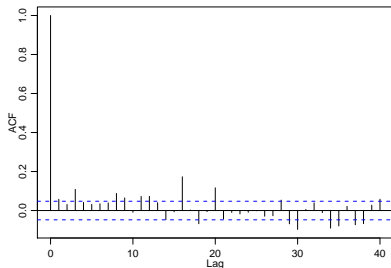


Evaluate the model selected in step two

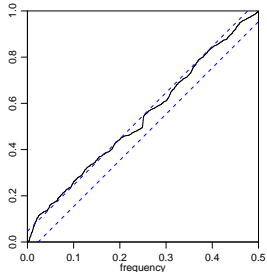
Inputs and residuals



ACF of residuals

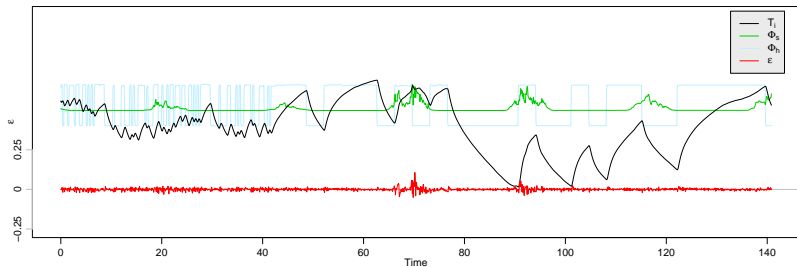


Cumulated periodogram

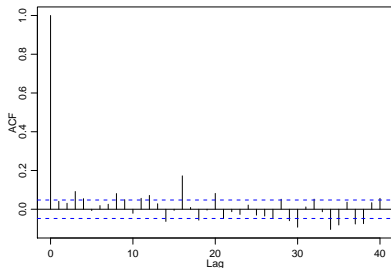


Evaluate the model selected in step three

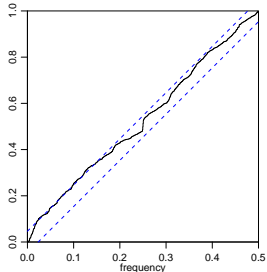
Inputs and residuals



ACF of residuals

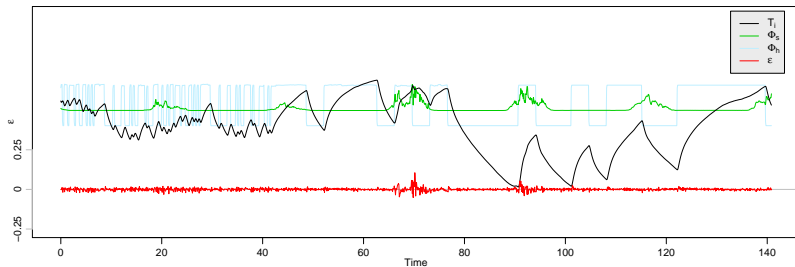


Cumulated periodogram

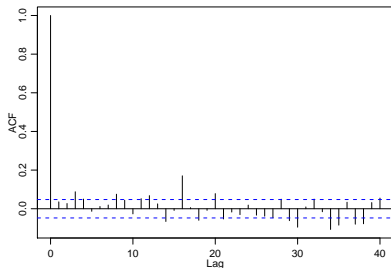


Evaluate the selected model in step four

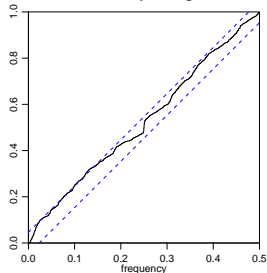
Inputs and residuals



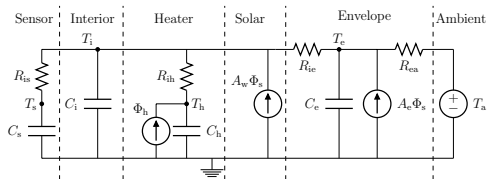
ACF of residuals



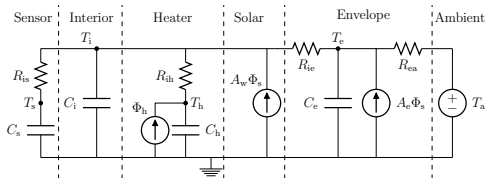
Cumulated periodogram



Selected model



Selected model



Estimated parameters

\hat{C}_i	0.0928	(kWh/°C)
\hat{C}_e	3.32	-
\hat{C}_h	0.889	-
\hat{C}_s	0.0549	-
\hat{R}_{ie}	0.897	(°ircC/kW)
\hat{R}_{ea}	4.38	-
\hat{R}_{ih}	0.146	-
\hat{R}_{is}	1.89	-
\hat{A}_w	5.75	(m ²)
\hat{A}_e	3.87	-

Estimated time constants

$\hat{\tau}_1$	0.0102	hours
$\hat{\tau}_2$	0.105	-
$\hat{\tau}_3$	0.788	-
$\hat{\tau}_4$	19.3	-

Conclusions

- Applied Grey-box modelling, where a combination of *prior physical knowledge* and *data-driven modelling* is utilized
- Using a forward selection procedure with likelihood-ratio tests a suitable physical model is found
- The ability of the selected models to describe the heat dynamics are evaluated with the ACF, CP, and time series plots

Identifiability

Identifiability

Model identifiability is important for estimation in general (less important for prediction, very important for parameter interpretation).

There are two aspects of identifiability:

- **Structural identifiability:** the parameters in the model can never be estimated due to the structure of the model. Depends only on the model.
- **Practical identifiability:** there is not enough information in the data available to estimate the parameters in the model. Depends both on the model and the data.

Structural identifiability

State space model (innovation form)

$$\begin{aligned}\frac{d\hat{X}(t)}{dt} &= A\hat{X}(t) + BU(t) + K\epsilon(t) \\ Y(t) &= C\hat{X}(t) + DU(t) + \epsilon(t)\end{aligned}$$

Apply the bilateral Laplace transformation (and after some voodoo)

$$\begin{aligned}Y(s) &= C(sI - A)^{-1}BU(s) + C(sI - A)^{-1}K\epsilon(s) + DU(s) + \epsilon(s) \\ &= \left(C(sI - A)^{-1}B + D\right)U(s) + \left(C(sI - A)^{-1}K + I\right)\epsilon(s)\end{aligned}$$

Focus on the input related transfer function

$$H_i(s) = C(sI - A)^{-1}B + D \quad (2)$$

Analyse the identifiability of an SDE model of a Wall

A lumped RC model of the wall

$$dT_w = \frac{1}{C_w} \left(\frac{T_a - T_w}{R_{aw}} + \frac{T_i - T_w}{R_{wi}} \right) dt + d\omega_1(t)$$

$$dT_i = \frac{1}{C_i} \left(\frac{T_w - T_i}{R_{wi}} \right) dt + d\omega_2(t)$$

$$y_{t_k} = T_{i_{t_k}} + \sigma_{t_k}$$

Transfer function

Apply equation ?? to obtain the input transfer function

$$H_{input}(s) = \frac{\frac{1}{C_i C_w R_{aw} R_{wi}}}{s^2 + \frac{R_{aw} C_i + C_i R_{wi} + R_{aw} C_w}{C_i C_w R_{aw} R_{wi}} \cdot s + \frac{1}{C_i C_w R_{aw} R_{wi}}}$$

Transfer function

Apply equation ?? to obtain the input transfer function

$$H_{input}(s) = \frac{\frac{1}{C_i C_w R_{aw} R_{wi}}}{s^2 + \frac{R_{aw} C_i + C_i R_{wi} + R_{aw} C_w}{C_i C_w R_{aw} R_{wi}} \cdot s + \frac{1}{C_i C_w R_{aw} R_{wi}}}$$

And compare it to

$$H(s) = \frac{b_0}{s^2 + a_1 \cdot s + a_0}$$

Transfer function

Apply equation ?? to obtain the input transfer function

$$H_{input}(s) = \frac{\frac{1}{C_i C_w R_{aw} R_{wi}}}{s^2 + \frac{R_{aw} C_i + C_i R_{wi} + R_{aw} C_w}{C_i C_w R_{aw} R_{wi}} \cdot s + \frac{1}{C_i C_w R_{aw} R_{wi}}}$$

And compare it to

$$H(s) = \frac{b_0}{s^2 + a_1 \cdot s + a_0}$$

Only two independent equations

$$a_0 = \frac{1}{C_i C_w R_{aw} R_{wi}}$$

$$a_1 = \frac{R_{aw} C_i + C_i R_{wi} + R_{aw} C_w}{C_i C_w R_{aw} R_{wi}}$$

Fit all four parameters?

Solve two equations for four parameters.

$$C_i = C_i$$

$$R_{wi} = R_{wi}$$

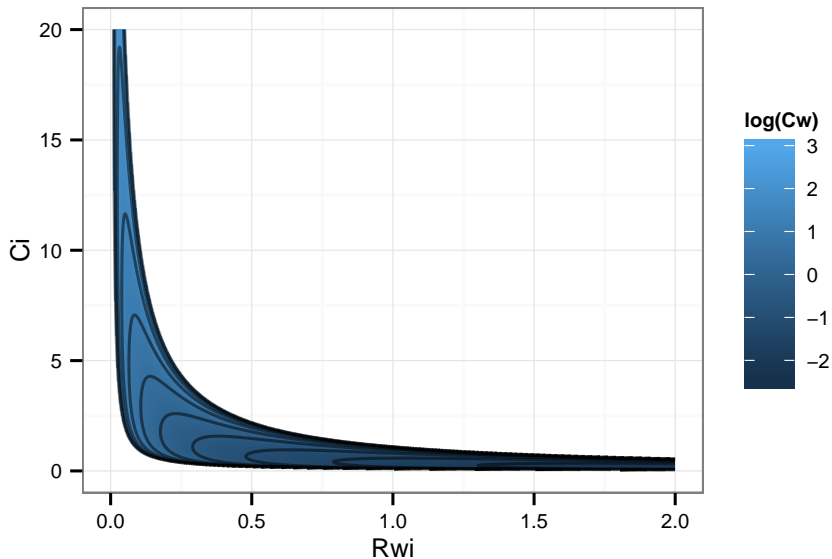
$$C_w = -\frac{C_i}{C_i^2 R_{wi}^2 a_0 - a_1 C_i R_{wi} + 1}$$

$$R_{aw} = -\frac{C_i^2 R_{wi}^2 a_0 - a_1 C_i R_{wi} + 1}{C_i^2 R_{wi} a_0}$$

Note: a_0 and a_1 are known when simulating data.

C_w is a function of other parameters

Below is the feasible C_w parameters: $C_w > 0$

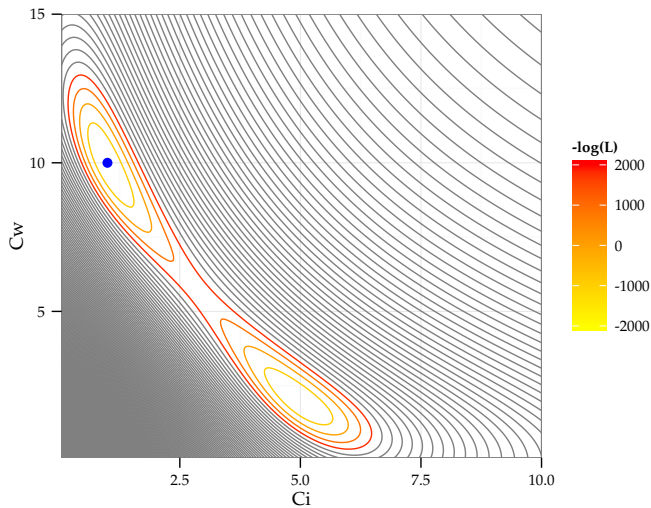


Estimate two parameters

We can estimate two.. So try fixing R_{wi} and R_{aw}

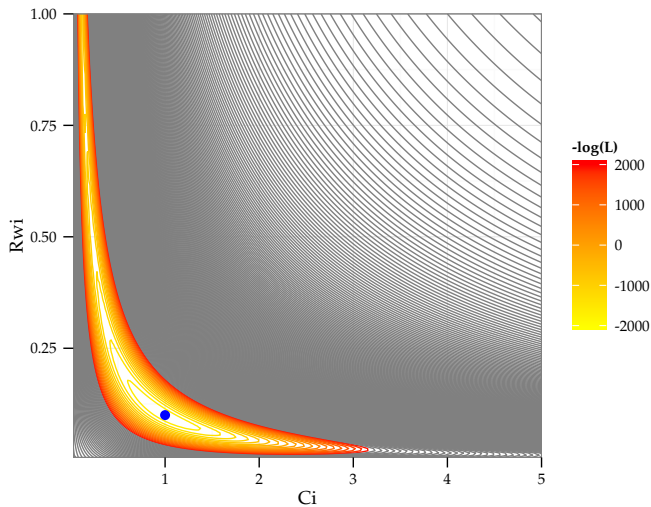
Estimate two parameters

We can estimate two.. So try fixing R_{wi} and R_{aw}



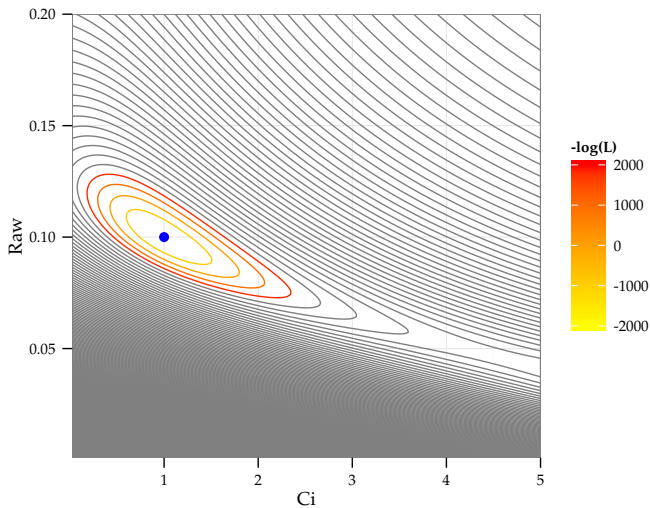
Estimate two parameters

We can estimate two.. So try fixing C_w and R_{aw}



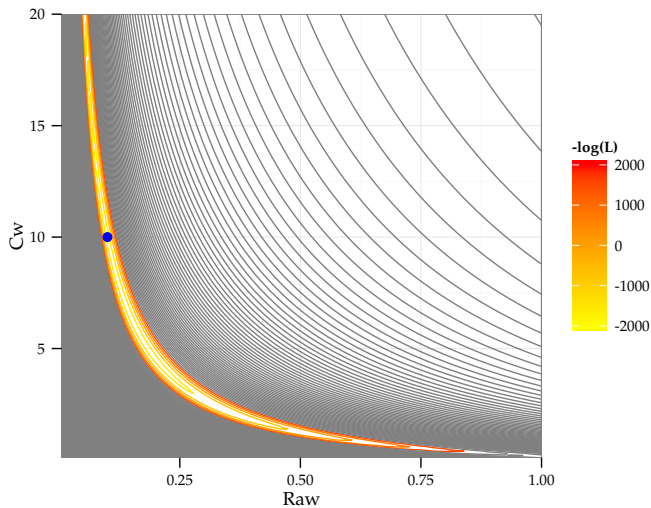
Estimate two parameters

We can estimate two.. So try fixing C_w and R_{wi}



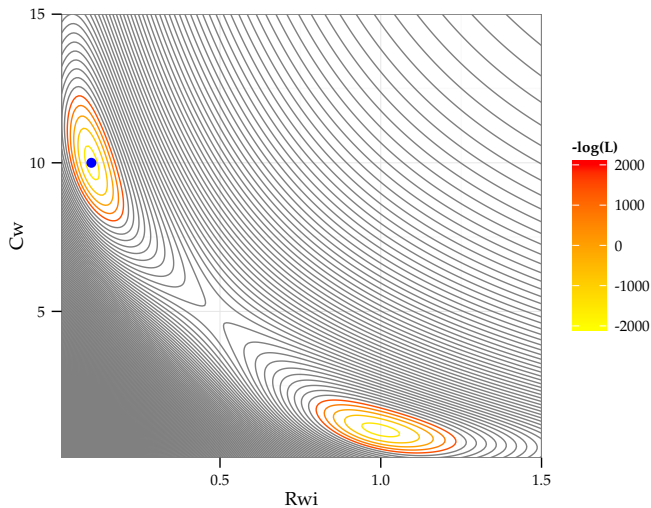
Estimate two parameters

We can estimate two.. So try fixing R_{twi} and C_i



Estimate two parameters

We can estimate two.. So try fixing R_{aw} and C_i



Estimate two parameters

We can estimate two.. So try fixing C_i and C_w

