

Time Series Analysis

Week 6 - ACF and PACF with a focus on model order selection

Peder Bacher

Department of Applied Mathematics and Computer Science
Technical University of Denmark

March 8, 2024

Week 6: Outline of the lecture

- ▶ Estimation of auto-covariance and -correlation, Sec. 6.2.1 (and the intro. to 6.2)
- ▶ Using the SACF and SPACF for model order selection Sec. 6.5
- ▶ Model validation, Sec. 6.6

Autocorrelation and Partial Autocorrelation

Autocorrelation

$$\rho(k) = \text{Cor}[Y_t, Y_{t+k}]$$

- ▶ Sample autocorrelation function (SACF): $\hat{\rho}(k) = r_k = C(k)/C(0)$
- ▶ For white noise and $k \neq 0$ it holds that $E[\hat{\rho}(k)] \simeq 0$ and $V[\hat{\rho}(k)] \simeq 1/N$, this gives the bounds $\pm 2/\sqrt{N}$ for deciding when it is not possible to distinguish a value from zero.
- ▶ R: `acf(x)`

Partial autocorrelation

$$\phi_{kk} = \text{Cor}[Y_t, Y_{t+k} | Y_{t+1}, \dots, Y_{t+k-1}]$$

- ▶ Sample partial autocorrelation function (SPACF): Use the Yule-Walker equations on $\hat{\rho}(k)$ (exactly as for the theoretical relations Eq.(5.81)) or as in next slide
- ▶ It turns out that $\pm 2/\sqrt{N}$ is also appropriate for deciding when the SPACF is zero
- ▶ R: `acf(x, type="partial")` or `pacf(x)`

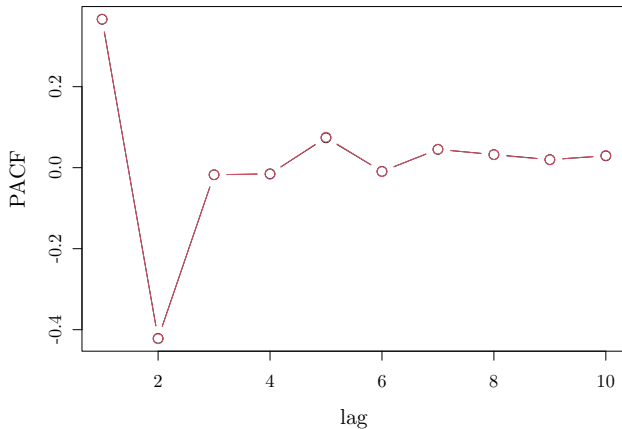
Autocorrelation and Partial Autocorrelation

```
# Example to show how the PACF is calculated
set.seed(972)
n <- 1000
x <- arima.sim(list(ar=c(0.5,-0.4)), n=n)
#acf(x)
#pacf(x)
library(onlineforecast)
D <- lagdf(c(x), 0:50)

# A way to calculate the PACF
lag.max <- 10
pacf1 <- numeric(lag.max)
# First, calculate it with the function
val <- pacf(x, lag.max, plot=FALSE)
# Then calc on our own
for(k in 1:lag.max){
  (frml <- pst("k0 ~ 1 + ",pst("k",1:k, collapse=" + ")))
  fit <- lm(frml, D)
  pacf1[k] <- fit$coef[pst("k",k)]
}
```

Autocorrelation and Partial Autocorrelation

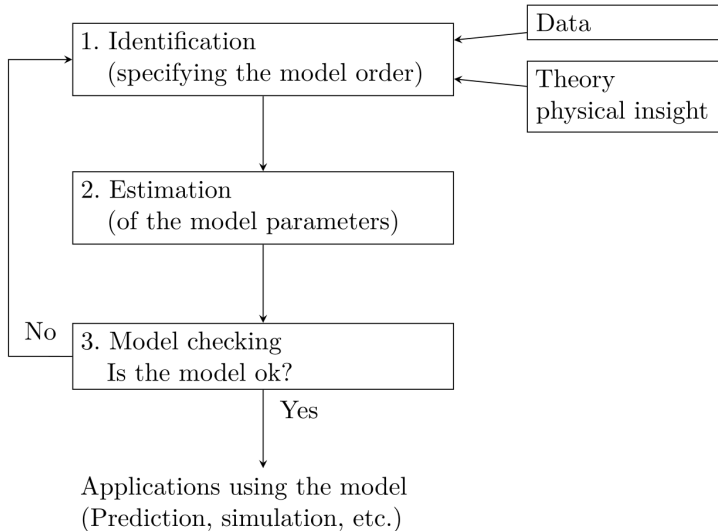
```
# It's very close!  
pacf1 - val$acf  
plot(val$acf, type="b", xlab="lag", ylab="PACF")  
lines(pacf1, type="b", col=2)
```



ARIMA models

- ▶ Today we will see how to identify ARIMA model orders
- ▶ Basically ARIMA(p, d, q) has
 - ▶ **Auto-Regression** order p
 - ▶ **Moving-Average** order q
 - ▶ **Integration** order d

Model building in general

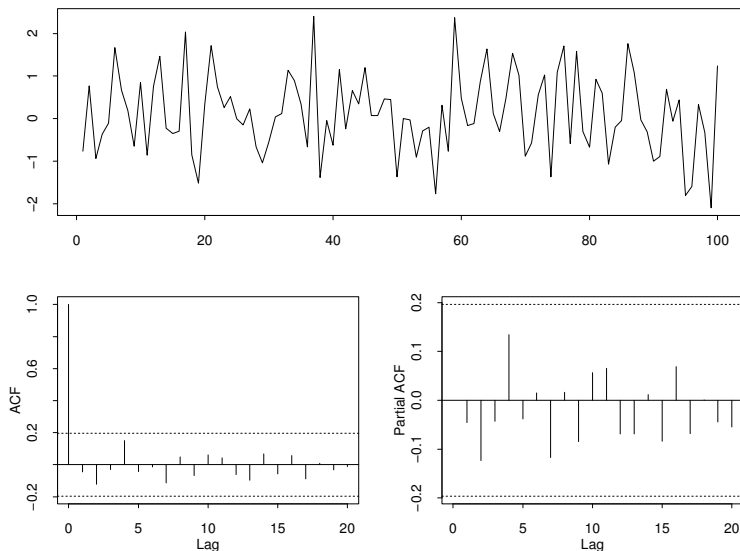


The golden table for ARMA identification

(Table 6.1)

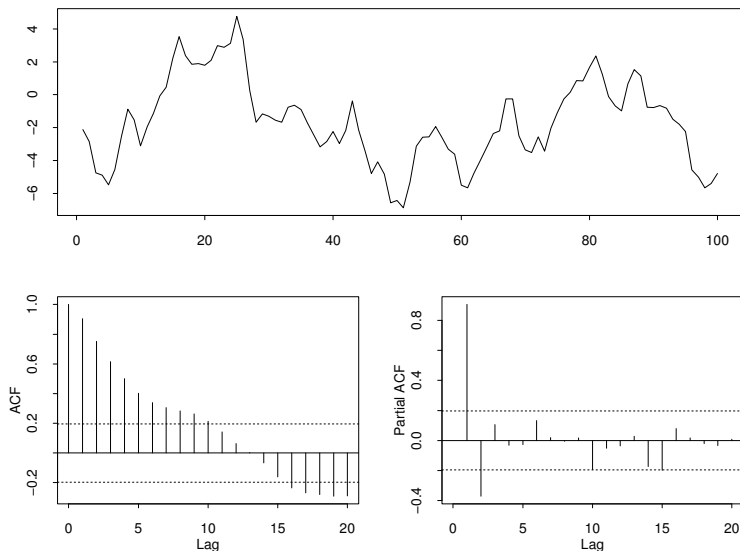
	ACF $\rho(k)$	PACF ϕ_{kk}
AR(p)	Damped exponential and/or sine functions	$\phi_{kk} = 0$ for $k > p$
MA(q)	$\rho(k) = 0$ for $k > q$	Dominated by damped exponential and or/sine functions
ARMA(p, q)	Damped exponential and/or sine functions after lag $\max(0, q - p)$	Dominated by damped exponential and/or sine functions after lag $\max(0, p - q)$

What would be an appropriate structure?



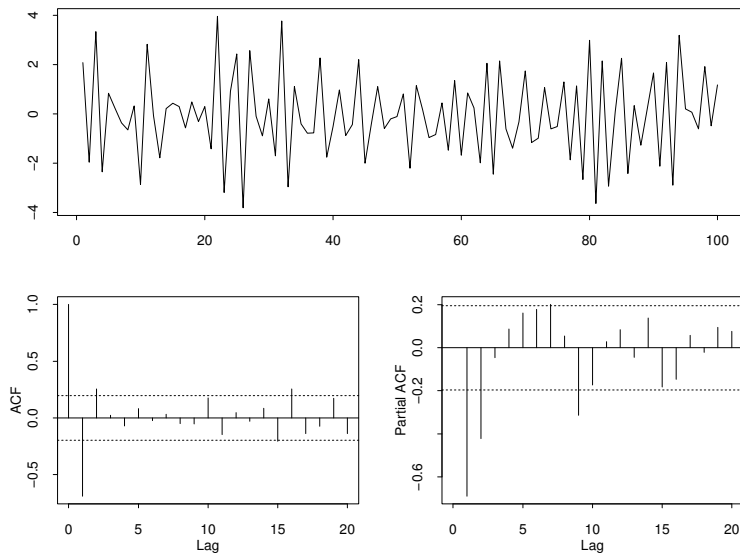
- 1: White noise 2: AR(1) 3: AR(2)
4: MA(1) 5: MA(2) 6: ARMA(1,1)

What would be an appropriate structure?



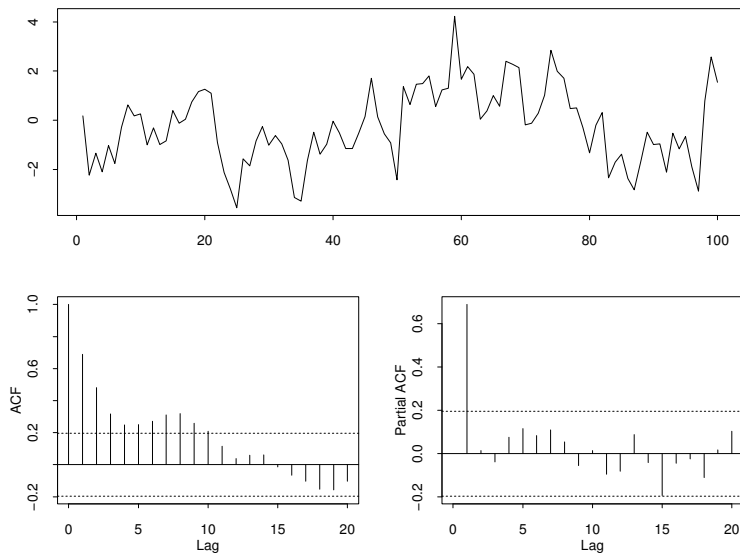
- 1: White noise 2: AR(1) 3: AR(2)
4: MA(1) 5: MA(2) 6: ARMA(1,1)

What would be an appropriate structure?



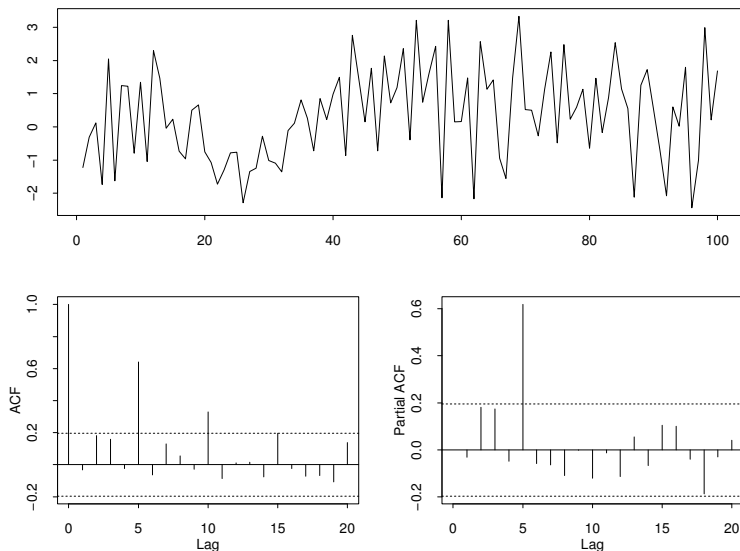
- 1: White noise 2: AR(1) 3: AR(2)
4: MA(1) 5: MA(2) 6: ARMA(1,1)

What would be an appropriate structure?



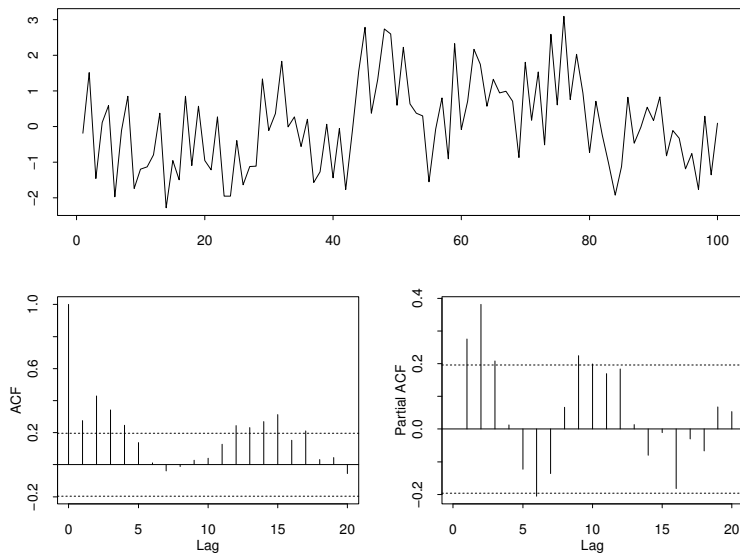
- 1: White noise 2: AR(1) 3: AR(2)
4: MA(1) 5: AR(8) 6: ARMA(1,1)

What would be an appropriate structure?



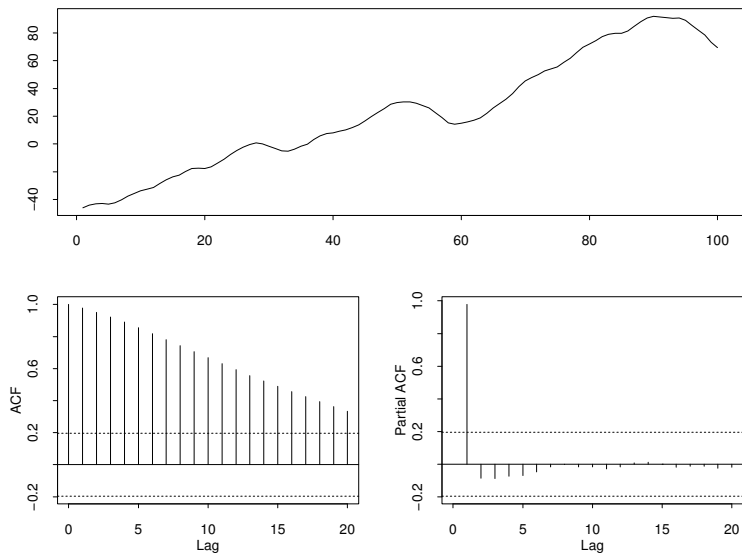
- 1: AR(1) 2: AR(5) 3: AR(0)x(2)₅
4: AR(0)x(1)₄ 5: AR(0)x(1)₅ 6: AR(10)

What would be an appropriate structure?



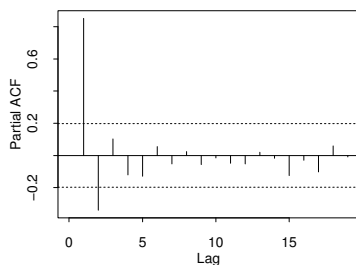
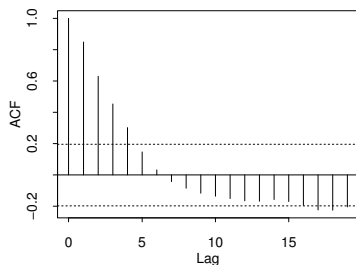
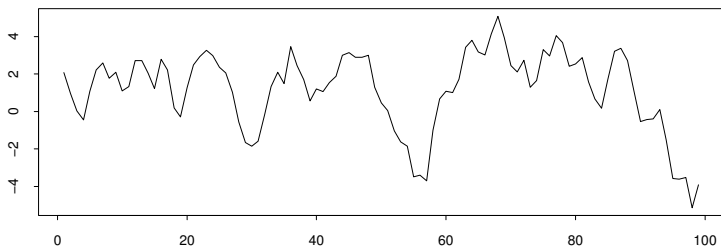
- 1: AR(4) 2: AR(5) 3: MA(4)
4: ARMA(1,1) 5: AR(2) 6: ARMA(2,2)

What would be an appropriate structure?



- | | | |
|--------------|-----------------|-----------------|
| 1: AR(3) | 2: AR(2) | 3: ARIMA(0,1,0) |
| 4: ARMA(1,1) | 5: ARIMA(2,1,0) | 6: ARIMA(0,2,2) |

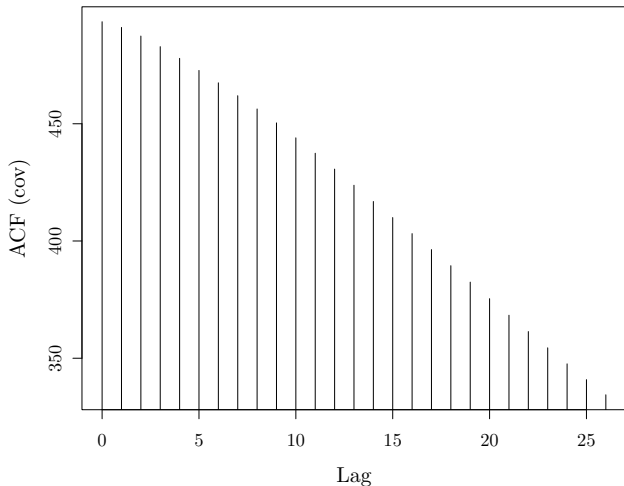
Same series; analysing $\nabla Y_t = (1 - B)Y_t = Y_t - Y_{t-1}$



- 1: AR(3) 2: AR(2) 3: ARIMA(0,1,0)
4: ARMA(1,1) 5: ARIMA(2,1,0) 6: ARIMA(0,2,2)

How does $C(k)$ behave for non-stationary series?

$$C(k) = \frac{1}{N} \sum_{t=1}^{N-|k|} (Y_t - \bar{Y})(Y_{t+|k|} - \bar{Y})$$

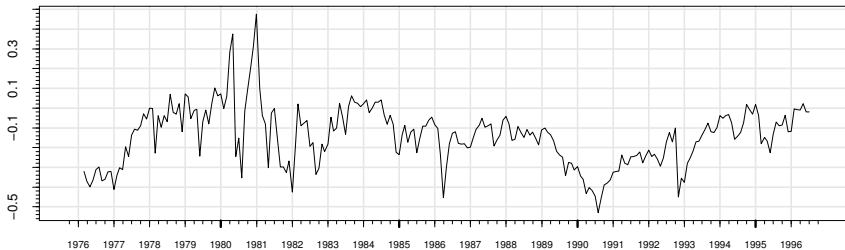


Identification of the order of differencing

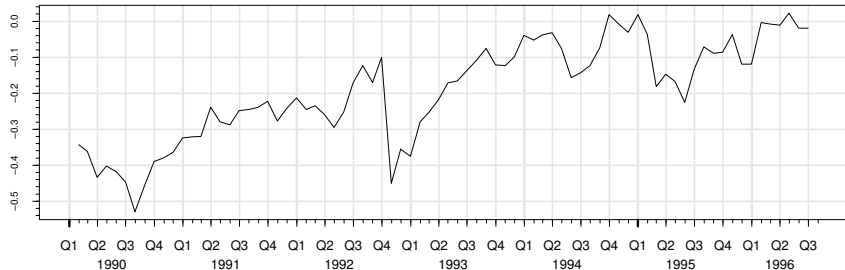
- ▶ Select the order of differencing d as the first order for which the autocorrelation decreases sufficiently fast towards 0
- ▶ In practice d is 0, 1, or maybe 2
- ▶ Sometimes a periodic difference is required, e.g. $Y_t - Y_{t-12}$
- ▶ Remember to consider the practical application. E.g. it may be that the system is stationary, but you measured over a too short period

Stationarity vs. length of measuring period

US/CA 30 day interest rate differential



US/CA 30 day interest rate differential



Selection of the Model Order

- ▶ The model order of an ARMA process model:
The number of parameters for the AR and MA part; (p, q) .
- ▶ The autocorrelation functions can be used - as we just did
- ▶ If that method fails to identify (p, q) because the process:
 - ▶ Is not a standard AR-proces.
 - ▶ Is not a standard MA-proces.
 - ▶ Is not a directly identifiable ARMA proces
- ▶ then try a small model and analyse the residuals
- ▶ and/or Consider transformations
Typically sqrt, log, square or inverse.

Iterative model building

1. (Identification step): Construct a model for your data:

$$\phi(B)Y_t = \theta(B)\varepsilon_t$$

2. (Estimation step): Estimate the coefficients $(\phi_1, \dots, \phi_p, \theta_1, \dots, \theta_q)$ and calculate the model residuals $\hat{\varepsilon}_{t|t-1}$
 3. (Model checking step):
 - ▶ Are the estimated coefficients significant?
 - ▶ Does $\hat{\varepsilon}_{t|t-1}$ resemble white noise?
 - ▶ If so, the model can be described by the ϕ and θ polynomials.
- ▶ If the model residuals do not resemble white noise, then what do they look like?

Iterative model building II

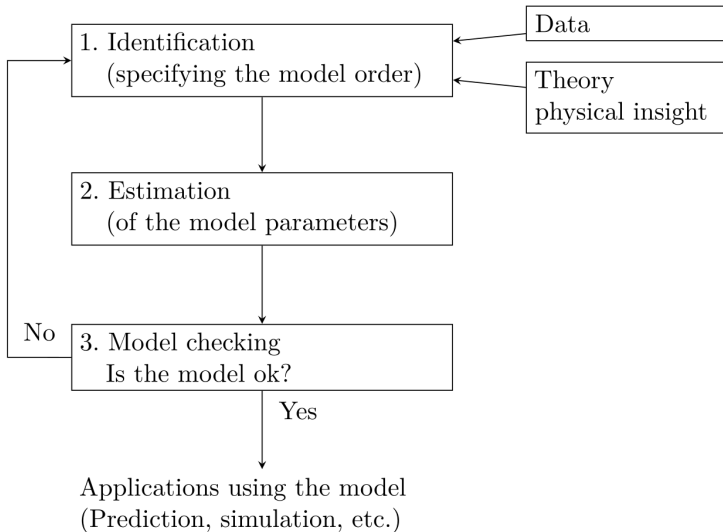
► $\hat{\varepsilon}_{t+k|t}$ will often have a simpler behavior than Y , if the original model $\phi(B)Y_t = \theta(B)\varepsilon_t$ captures the essential terms of Y 's behavior.

1. Construct an ARMA description for $\hat{\varepsilon}_{t|t-1}$: $\phi^*(B)\varepsilon_t = \theta^*(B)\varepsilon_t^*$.
2. Insert $\varepsilon_t = \phi^{*-1}(B)\theta^*(B)\varepsilon_t^*$ into the original model to obtain the model

$$\phi^*(B)\phi(B)Y_t = \theta(B)\theta^*(B)\varepsilon_t^*$$

3. Estimate the parameters in the model above with coefficients in $\phi^* \cdot \phi$, $\theta \cdot \theta^*$ varying freely, and proceed to model check.

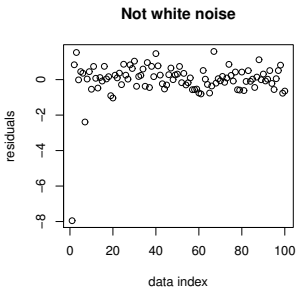
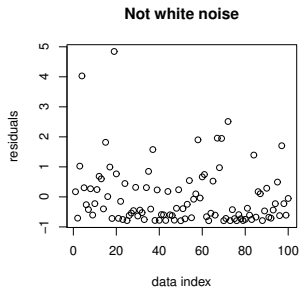
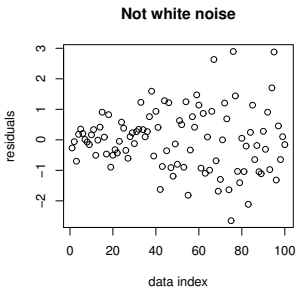
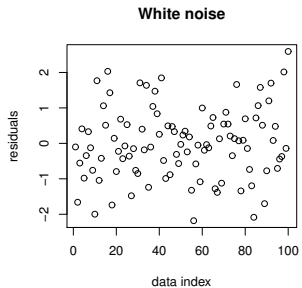
Model building in general



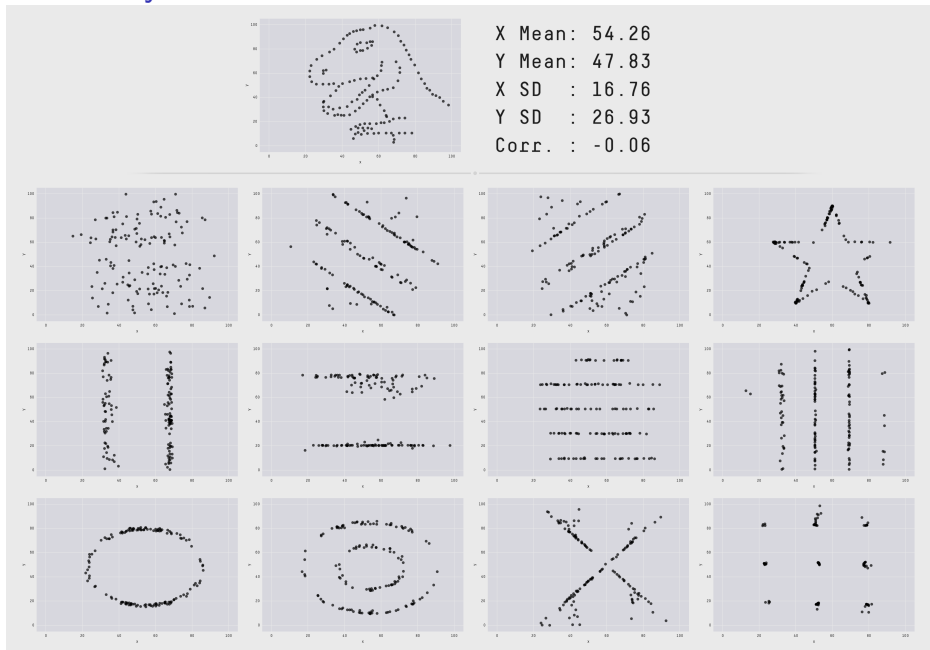
Residual Analysis

- ▶ The order of the model is the minimum order for which the model errors resemble white noise.
- ▶ How can we check that the model errors resemble white noise?
- ▶ First and most important - plot the data.

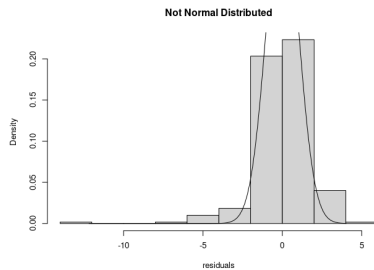
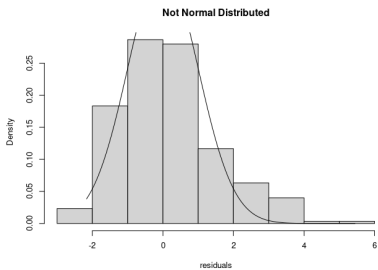
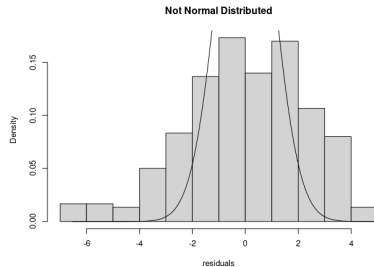
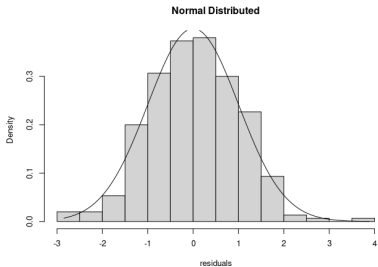
Residual analysis – Plot the data



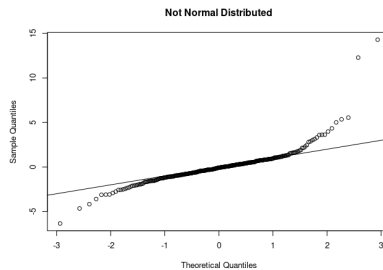
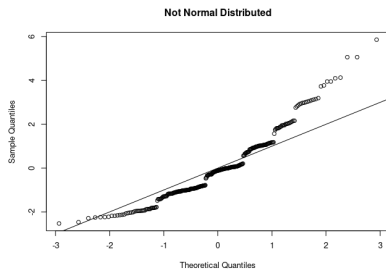
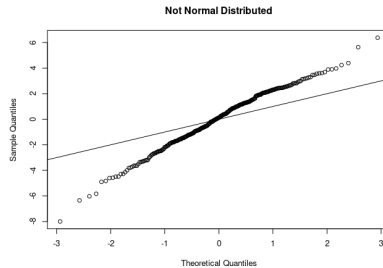
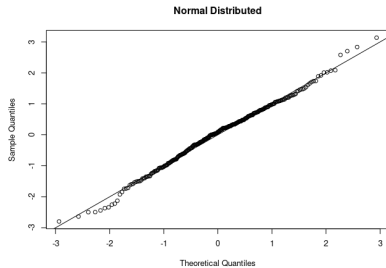
Residual analysis – Plot the data II



Residual analysis – Plot the data III

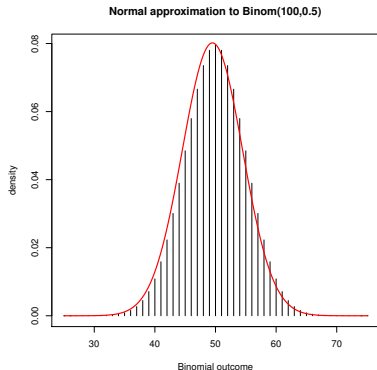


Residual analysis – Plot the data IV

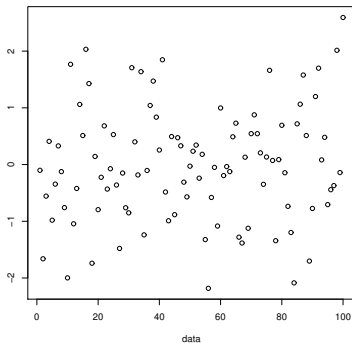


Residual analysis – sign test

- ▶ If (ε_t) is white noise, the probability that a new value has a different sign than the previous is $\frac{1}{2}$.
- ▶ Number of sign changes: $\text{Binom}(N - 1, \frac{1}{2})$.
- ▶ Approx. normal distribution; $N((N - 1)/2, (N - 1)/4)$:



Residual analysis – sign test II



- ▶ 95% confidence interval for sign changes within 100 white noise residuals: $[40; 59]$. Actual sign changes from the 100 data: 47.

Residual analysis – sign test III

Sign tests detects both asymmetry and correlation.

- ▶ Too few may indicate positive one-step correlation;
- ▶ Too many may indicate negative one-step correlation;
- ▶ Too few or too many may indicate that $P(\text{being above the mean}) \neq \frac{1}{2}$ with no correlation.

Residual analysis - other tests

- ▶ There is a bunch of other tests out there.
- ▶ You are welcome to use them in assignments, as long as you are sure that you understand them.

Residual analysis – summary

- ▶ Plot $\hat{\varepsilon}_{t|t-1}$; do the residuals look stationary? Do they need a transformation?
- ▶ Plot estimated ACF and PACF, if there are significant lags, then can we use them to extend the model with an ARMA-structure?
- ▶ Plot histogram and/or qq-plot to see whether residuals are normal distributed, if not, then consider a transformation.
- ▶ Perform a couple of statistical tests to get some quantitative measures of whether your residuals are alright.
- ▶ Finally, see whether parameters are significant and if not, remove them (you do not need to redo residuals analysis after this).

Information criteria

When considering multiple non-nested candidate models, information criteria can be used:

- ▶ Select the model which minimizes some information criterion.
- ▶ Akaike's Information Criterion:
$$AIC = -2 \log(L(Y_N; \hat{\theta}, \hat{\sigma}_\varepsilon^2)) + 2n_{\text{par}}$$
- ▶ Bayesian Information Criterion (preferred):
$$BIC = -2 \log(L(Y_N; \hat{\theta}, \hat{\sigma}_\varepsilon^2)) + \log(N)n_{\text{par}}$$
- ▶ AIC is most commonly used, but BIC yields a consistent estimate of the model order.

Cross validation

Cross-validation is possible but slightly less efficient and cumbersome for time series analysis than for other kinds of data.

- ▶ If we use future measurements we are cheating!
- ▶ Thus, it is only possible to split data by having first part be for training, and last part testing.



- ▶ So we must gradually move the part used for training forward in time, it's called "rolling horizon" cross-validation
- ▶ Mainly used for forecasting applications
- ▶ Remember a burn-in period and then step forward from there